

# 概率图模型理论及应用

Theory and Applications of Probabilistic Graphical Models  
(Lesson 10 - sampling)

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# 课前摘要

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abs\_lesson10\_Sampling\_章雅婷.



abs\_lesson10\_Sampling\_贾浩歌.



abs\_lesson10\_sampling\_王腾蛟.

# 课程章节

## ❖ 第一章 引言 (1)

## ❖ 第二章 图模型的表示理论 (2)

- Semantics (DGM, UGM)
- HMM, CRF

## ❖ 第三章 图模型的推理理论 (6)

- 精确推理: **variable-elimination**, **cluster-tree**, **triangulate**
- 连续变量: **Kalman** 树消除算法 (Cluster-tree elimination)  
理论上总是成立
- 采样近似: **sampling**
- 变分近似: **variational**

## ❖ 第四章 图模型的学习理论 (3)

- 参数学习: **maxlikelihoodEstimate**, **RFLearning**, **BayesEstimate**
- 结构学习: **StructureLearning**

## ❖ 第五章 一个综合例子 (1)

# 2 cases for approximate inference

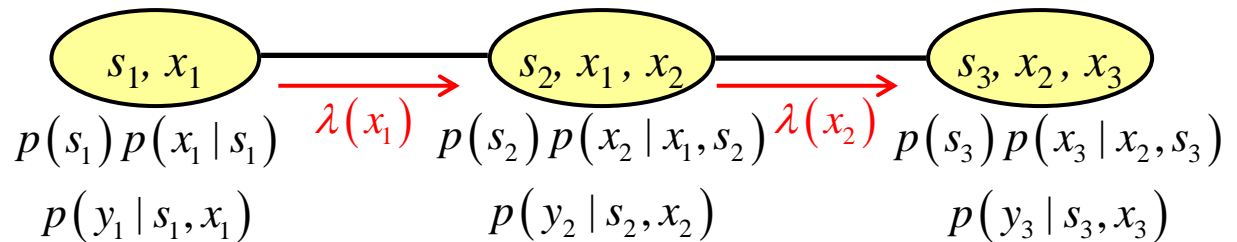
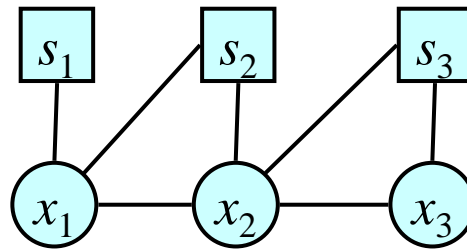
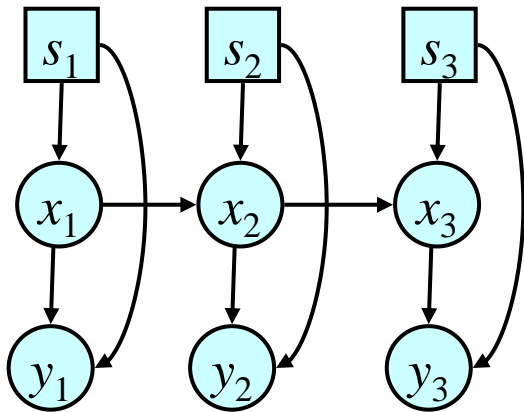
根源：消息计算实际上无法进行

$$\lambda_{C_1 \rightarrow C_2} \stackrel{= \Downarrow}{\text{sep}(C_1, C_2)} \{ \prod \dots \}$$

**Case 1: containing continuous-var, nonlinear though low induced-width**

Exact only when *either* LG, or CG with continuous-var elim first

例如：switching KF  $p(x_t | y_{1:t}) = ?$



表示了  $M$  个高斯（假设  $i$  有  $M$  个可能取值）

$$\sum_i \phi(x, i | g_i, h_i, K_i) = \phi(x | \{g_i, h_i, K_i\}_{1 \leq i \leq M})$$

消除变量  $i$ ，得到  $M$  个高斯的混合

# 2 cases for approximate inference

根源：消息计算实际上无法进行

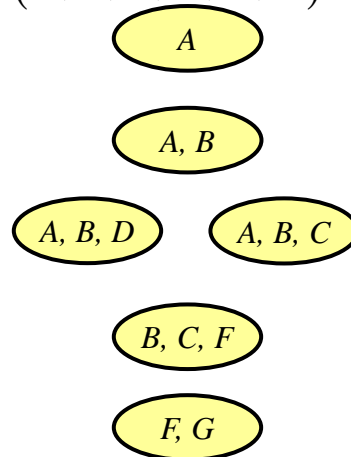
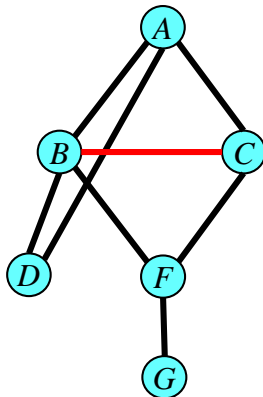
## Case 2: High induced-width $w(G)$ , $G$ : 连乘积的素图

在均是离散变量情形下，推理计算的复杂度与连乘积的素图表示的诱导宽度成指数关系。

最优树分解中最大的cluster

- 图 $G$ 的诱导宽度  $w(G) =$  图 $G$  在所有可能排序下的诱导宽度的最小者  
在最优排序下的诱导宽度
- 图 $G$ 在排序 $d$ 下的诱导宽度  $w(G_d)$ 
  - = 在排序 $d$ 下，消除结点过程中，结点的邻居数目的最大者
  - = 在排序 $d$ 下，消除式三角化过程中，结点的最大消除集的规模 - 1

排序 $d = (A, B, C, D, F, G)$ 下消除集



# Example: Image de-noising

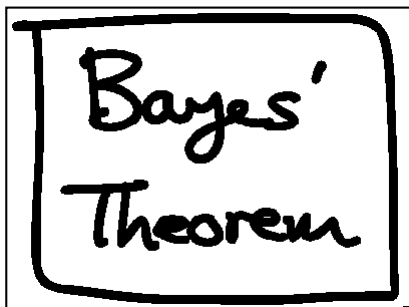
## ❖ 带噪图像的概率模型

- $x_i$ : 原始干净图像像素  $\in \{-1, 1\}$
- $y_i$ : 带噪观测图像像素  $\in \{-1, 1\}$
- 相邻象素  $x_i, x_j$  取值的随机规律
- $y_i$  与  $x_i$  取值的随机规律

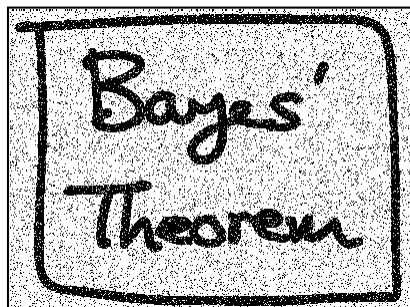
$$p(x, y) \propto \exp \left\{ \beta \sum_{i-j} x_i x_j + \gamma \sum_i x_i y_i \right\} \quad \beta > 0, \gamma > 0$$

## ❖ 给定带噪观测图像 $y$ , 恢复原始干净图像 $x$

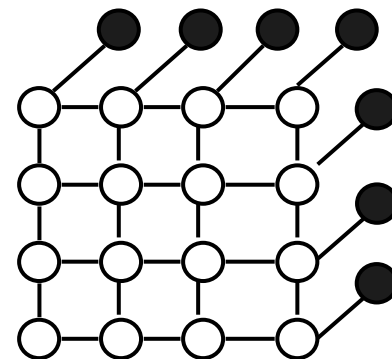
$$\max_{x_i} p(x_i | y)$$



original clean



noisy observation

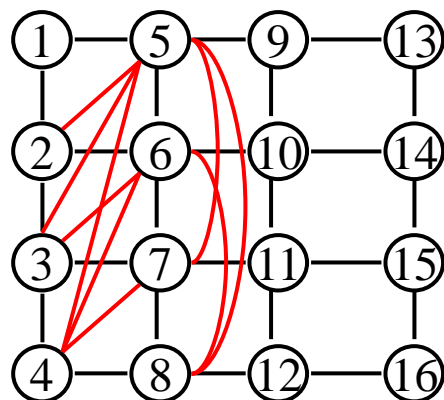


$$\phi(x_i, x_j)^{i-j} = \begin{array}{c|cc} & -1 & 1 \\ \hline -1 & e^\beta & e^{-\beta} \\ 1 & e^{-\beta} & e^\beta \end{array}$$

$$\phi(x_i, y_i) = \begin{array}{c|cc} & -1 & 1 \\ \hline -1 & e^\gamma & e^{-\gamma} \\ 1 & e^{-\gamma} & e^\gamma \end{array}$$

连乘积形式联合分布  $p(x, y) \propto \exp \left\{ \beta \sum_{i-j} x_i x_j + \gamma \sum_i x_i y_i \right\} \quad \beta > 0, \gamma > 0$

连乘积的素图表示  $G$



induced-width  $w(G) = ?$

找出最优变量排序，保证尽可能小的消除集

- |           |              |
|-----------|--------------|
| 1,2,5     | 5,6,7,8,9    |
| 2,3,5,6   | 6,7,8,9,10   |
| 3,4,5,6,7 | 7,8,9,10,11  |
| 4,5,6,7,8 | 8,9,10,11,12 |

# Introduction - sampling

## ❖ 统计模拟方法—Monte Carlo 方法：非数值方法计算积分

- 欲计算积分  $E_{p(x)}[\phi(x)] = \int \phi(x) p(x) dx$

- 随机生成  $R$  个服从分布  $p(x)$  的独立样本  $x^{(r)}$ ，对上述积分近似计算为

$$\int \phi(x) p(x) dx \approx \frac{1}{R} \sum_{r=1}^R \phi(x^{(r)})$$

概率计算可看成是示性函数的期望

$$p(\bar{x}) = \int 1(x = \bar{x}) p(x) dx \approx \frac{1}{R} \sum_{r=1}^R 1(x^{(r)} = \bar{x})$$

采样近似推理： $p(x / E)$ 无法直接计算，获取后验分布  $p(x / E)$  的样本？

用概率分布  $p(x / E)$  的样本去表示这个分布



# Sampling approximation

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**MCMC: MH sampling, Gibbs sampling**

图像去噪

**Case 2: High induced-width  $w(G)$**

**Importance sampling, Particle filtering**

视频物体跟踪

**Case 1: containing continuous-var, nonlinear**

# MCMC example: Metropolis–Hastings algorithm

We want to Draw samples from the target distribution  $p(x)$  ?

Solution: Construct a Markov chain that has  $p(x)$  as the stationary distribution.

1. Random initialize  $x_0$

2. For  $t = 1, \dots$

Generates  $x^*$  from proposal/transition kernel  $q(x^*|x_{t-1})$ ,

Accept  $x_t = x^*$  with probability  $\min\left\{1, \frac{p(x^*)q(x_{t-1}|x^*)}{p(x_{t-1})q(x^*|x_{t-1})}\right\}$ ,

otherwise set  $x_t = x_{t-1}$

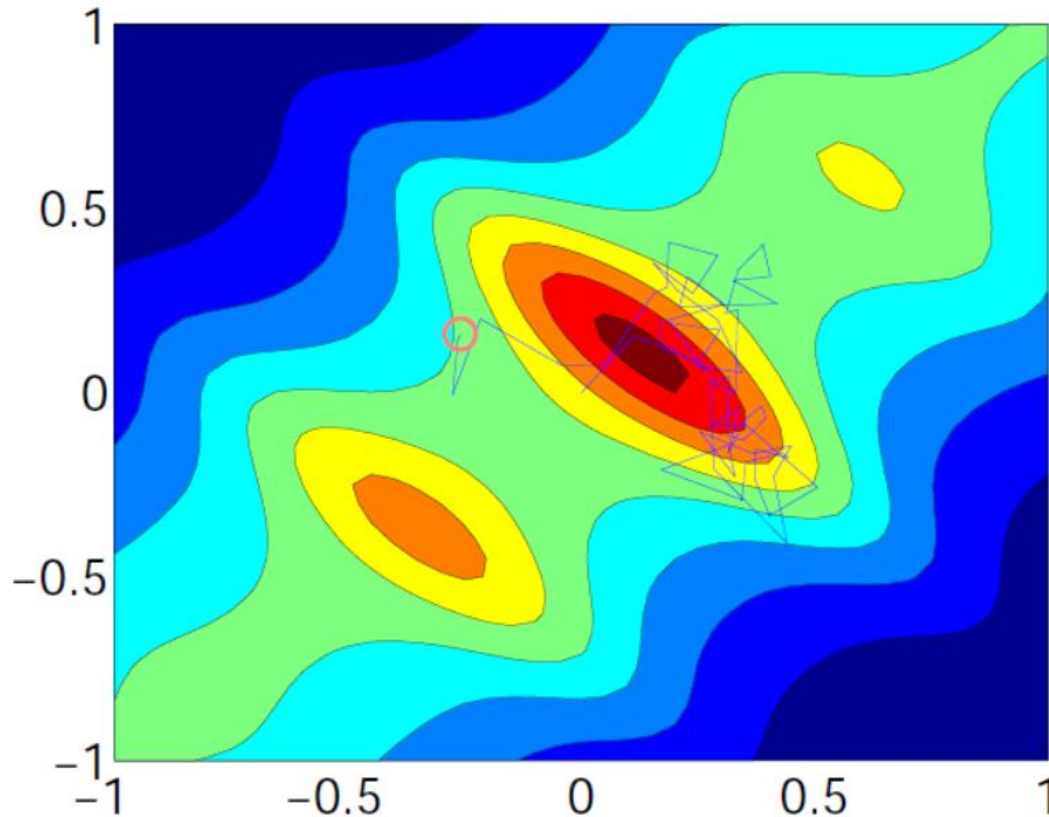
Burn-in : first few samples are discarded.

- For an irreducible & ergodic Markov chain, there exist a stationary distribution  $\pi$  (i.e. satisfying equation  $\pi = \pi P$ ).
- A sufficient condition: satisfy the detailed balance equation

$$\pi_i P_{ij} = \pi_j P_{ji}$$

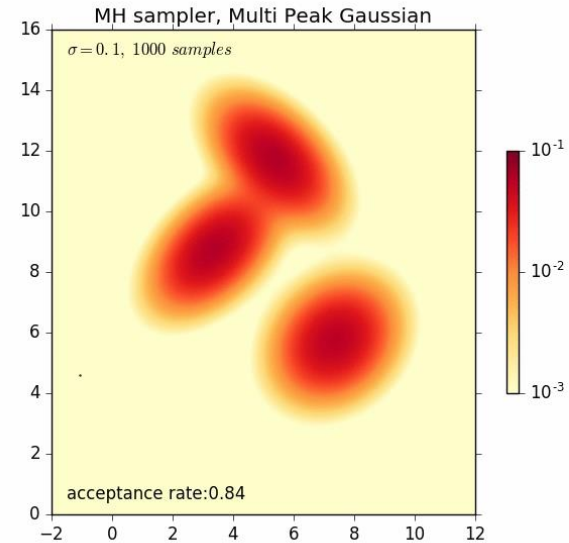
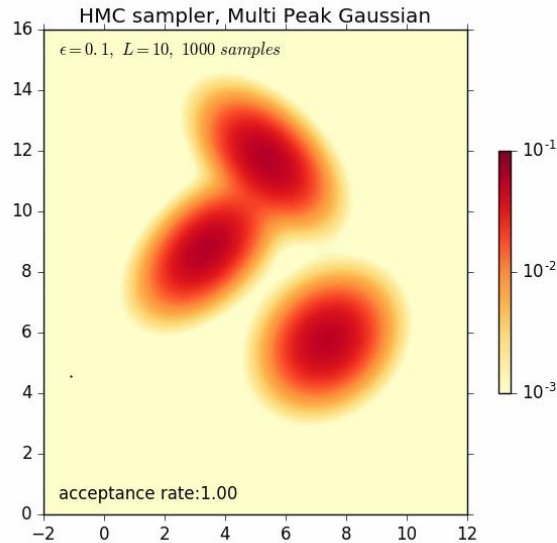
# Metropolis–Hastings example

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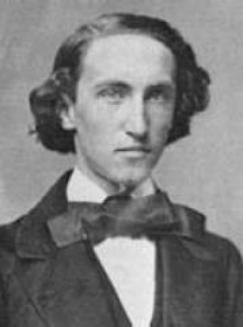


$$\text{e.g. } q(x^* | x_{t-1}) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x^* - x_{t-1})^2}{2\sigma^2}}$$

# Hamiltonian Monte Carlo (HMC) vs MH



GIF demo for sampling from a 2-dim 3-component GMM



J.W. Gibbs, 1839-1903

# Gibbs sampling

❖ 背后原理：Markov Chain Monte Carlo (MCMC)

❖ 用于多维分布的采样

❖ 考虑对一个多维分布  $p(\mathbf{z}) = p(z_1, \dots, z_M)$  的采样：

1. *Initialize*  $\{z_1^{(1)}, z_2^{(1)}, z_3^{(1)}, z_4^{(1)}\}$

2. A *sweep* generates a sample of  $\mathbf{z}^{(t)} = \{z_1^{(t)}, z_2^{(t)}, z_3^{(t)}, z_4^{(t)}\}$

每次固定其它分量的当前采样值，从一个分量的条件分布采样

– *Sample*  $z_1^{(2)} \sim p(z_1 | z_2^{(1)}, z_3^{(1)}, z_4^{(1)})$

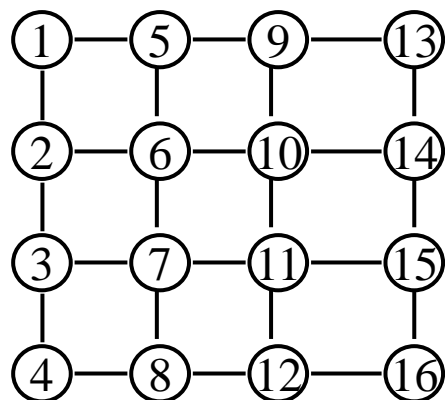
– *Sample*  $z_2^{(2)} \sim p(z_2 | z_1^{(2)}, z_3^{(1)}, z_4^{(1)})$

– *Sample*  $z_3^{(2)} \sim p(z_3 | z_1^{(2)}, z_2^{(2)}, z_4^{(1)})$

– *Sample*  $z_4^{(2)} \sim p(z_4 | z_1^{(2)}, z_2^{(2)}, z_3^{(2)})$

连乘积形式联合分布  $p(x, y) \propto \exp \left\{ \beta \sum_{i-j} x_i x_j + \gamma \sum_i x_i y_i \right\} \quad \beta > 0, \lambda > 0$

连乘积的素图表示



❖ 给定带噪观测图像  $y$ , 恢复原始干净图像  $x \quad \max_{x_i} p(x_i | y)$

1. *Initialize*  $\{x_1^{(1)}, \dots, x_{16}^{(1)}\}$

sample from  $p(x_1 | x_{-1}, y)$

sample from  $p(x_2 | x_{-2}, y)$

⋮

sample from  $p(x_{16} | x_{-16}, y)$

$\rightarrow \{x_1^{(2)}, \dots, x_{16}^{(2)}\} \rightarrow \{x_1^{(3)}, \dots, x_{16}^{(3)}\}$

# Sampling approximation

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## Gibbs sampling

图像去噪

**Case 2: High induced-width  $w(G)$**

## Importance sampling, Particle filtering

视频物体跟踪

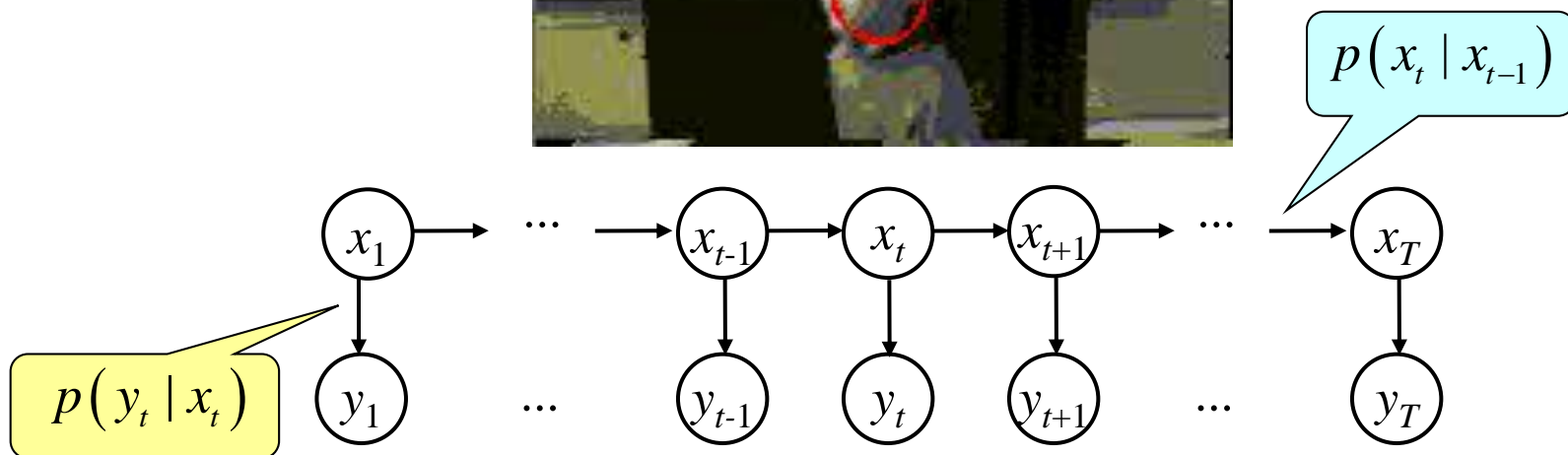
**Case 1: containing continuous-var, nonlinear**

# 视频目标跟踪 $\rightarrow$ 轮廓跟踪 求: $p(x_t | y_{1:t}) = ?$

- ❖ 观测量  $y_1, y_2, \dots$  为视频图像序列(帧序列)
- ❖ 状态量  $x_1, x_2, \dots$  为所关心物体的位置(轮廓线)



$$x_t = \begin{pmatrix} x_t^1 \\ x_t^2 \\ x_t^3 \\ \vdots \\ x_t^K \end{pmatrix}$$





# 轮廓跟踪—状态转移 $p(x_t | x_{t-1})$

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## ❖ 领域随机扰动

$$x_t = x_{t-1} + \omega_{t-1} \quad \omega_{t-1} \sim N(0, Q)$$

## ❖ 匀速运动的一阶模型

$$\begin{pmatrix} x_t \\ \dot{x}_t \end{pmatrix} = \begin{pmatrix} 1 & \Delta \\ 0 & 1 \end{pmatrix} \begin{pmatrix} x_{t-1} \\ \dot{x}_{t-1} \end{pmatrix} + \begin{pmatrix} 1 & 0 \\ 0 & 0 \end{pmatrix} \begin{pmatrix} \omega_{t-1,1} \\ \omega_{t-1,2} \end{pmatrix}$$

# 轮廓跟踪—观测模型

观测模型  $p(y_t | x_t)$  表示：给定图像  $y_t$  下，出现特定轮廓线  $x_t$  的似然值

沿上次跟踪结果法线的边缘特征

轮廓线  
contour

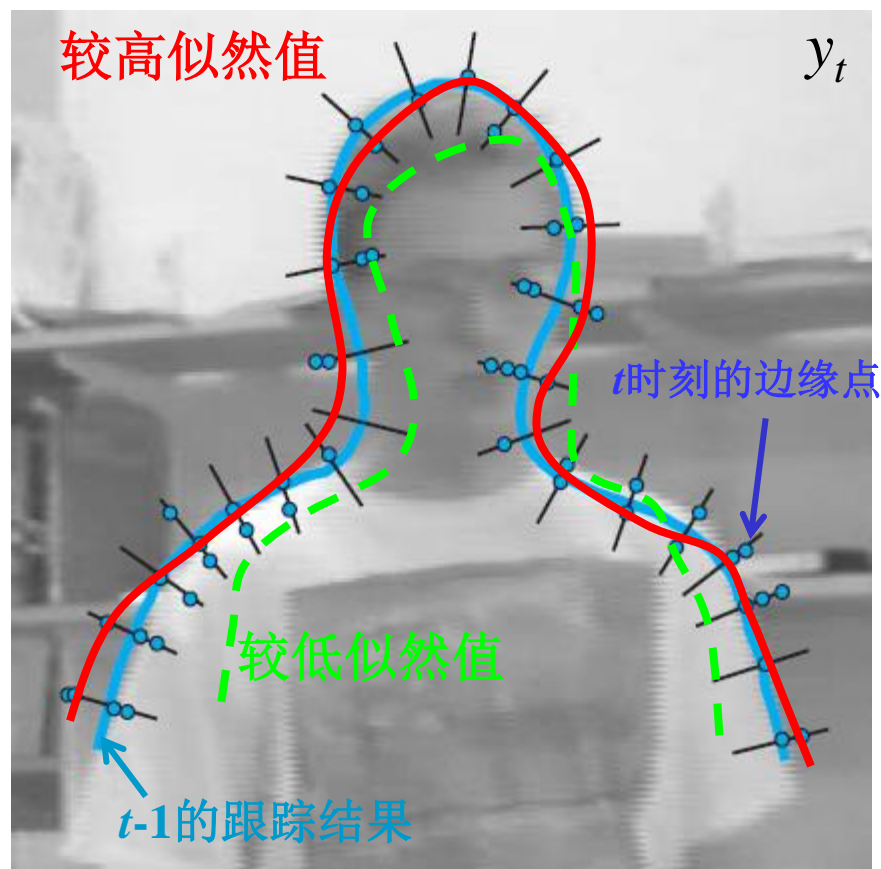
$$\begin{pmatrix} y_t^1 \\ y_t^2 \\ y_t^3 \\ \vdots \\ y_t^K \end{pmatrix}$$

沿着上次轮廓线的第  $k$  条法线上的边缘点

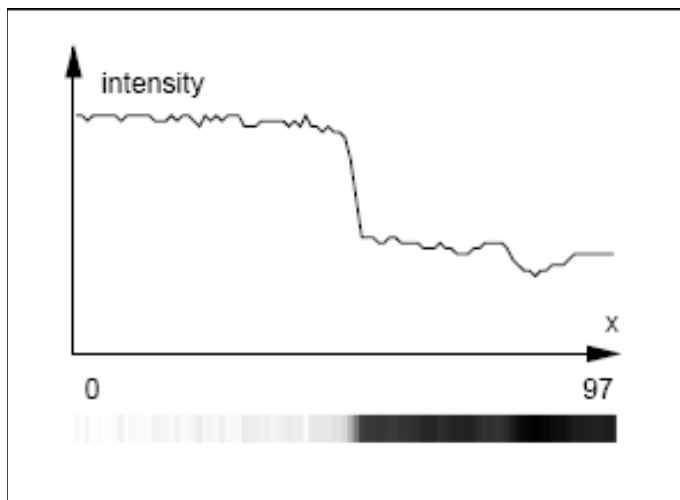
$$\begin{pmatrix} x_t^1 \\ x_t^2 \\ x_t^3 \\ \vdots \\ x_t^K \end{pmatrix}$$

轮廓线的  $K$  个支点—轮廓点

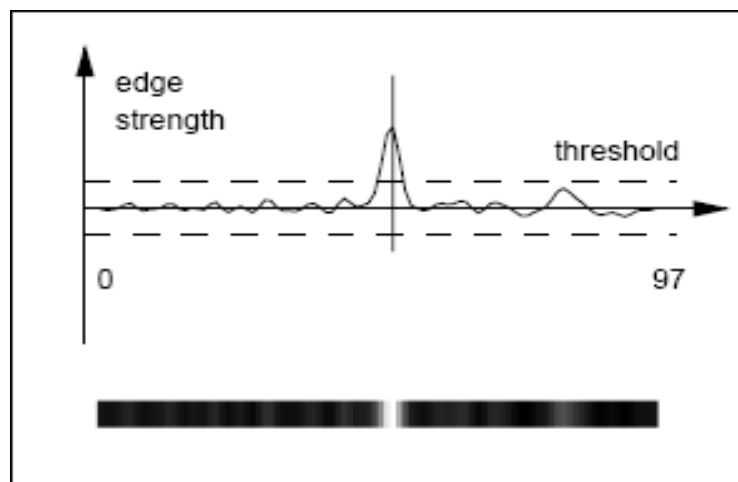
$$p(y_t | x_t) = \prod_k p(y_t^k | x_t^k)$$



# 轮廓跟踪—边缘特征

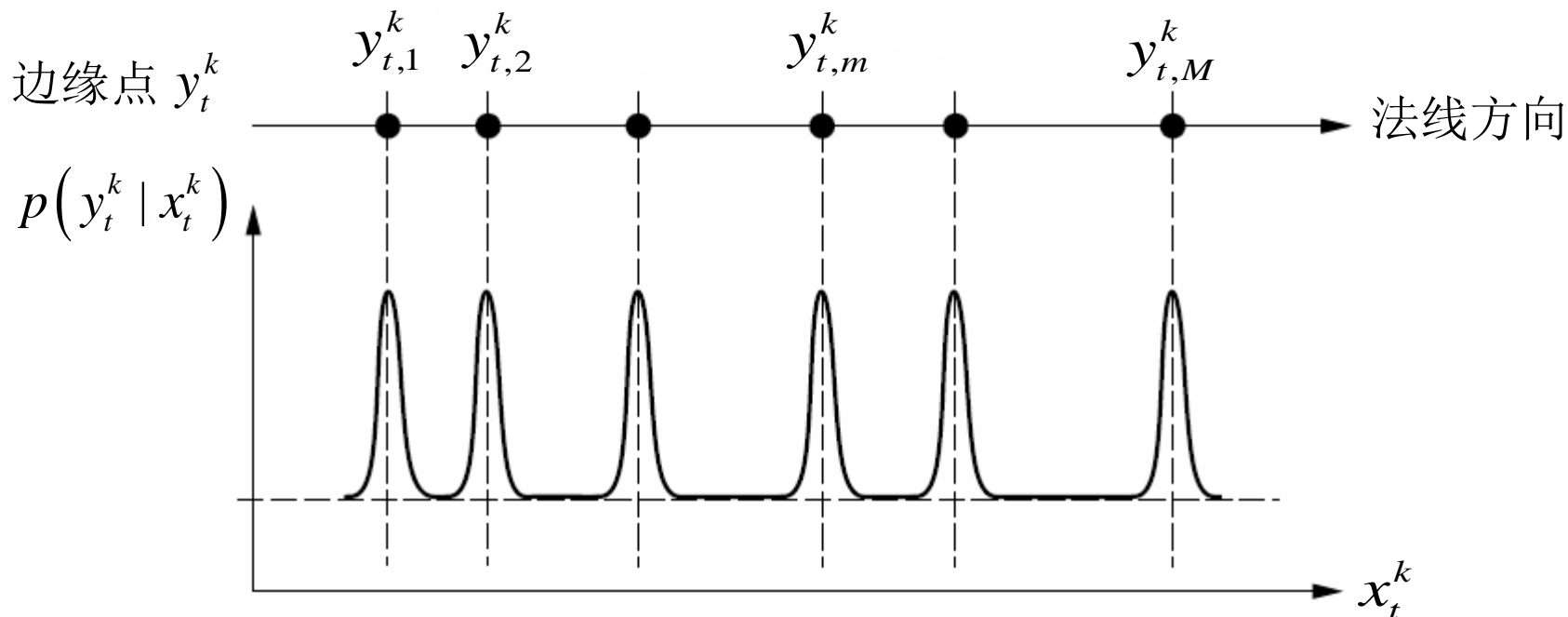


轮廓点的法线方向上灰度值



轮廓点的法线方向上灰度的梯度

# 轮廓跟踪—观测模型



$$p(y_t^k | x_t^k) = \frac{1}{M} \sum_{m=1}^M \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{\|y_{t,m}^k - x_t^k\|^2}{2\sigma^2}\right\}$$

轮廓线与第 $k$ 条法线的交点位置—轮廓点

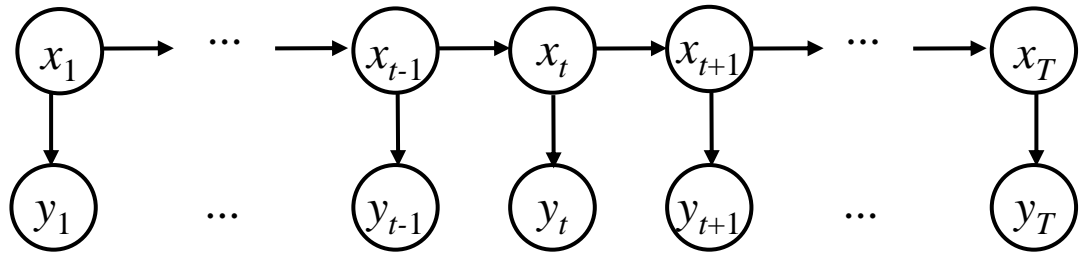
$$p(y_t | x_t) = \prod_{k=1}^K p(y_t^k | x_t^k)$$

$p(y_t | x_t)$  是  $M^K$  个典范函数的混合

**多峰**  
**非高斯**



$$p(x_t | y_{1:t}) = ?$$



测量更新

$$p(x_t | y_{1:t}) = \frac{p(y_{1:t-1})}{p(y_{1:t})} p(y_t | x_t) p(x_t | y_{1:t-1})$$

时间更新

$$p(x_{t+1} | y_{1:t}) = \int p(x_{t+1} | x_t) p(x_t | y_{1:t}) dx_t$$

特别地，**Kalman Filter** :

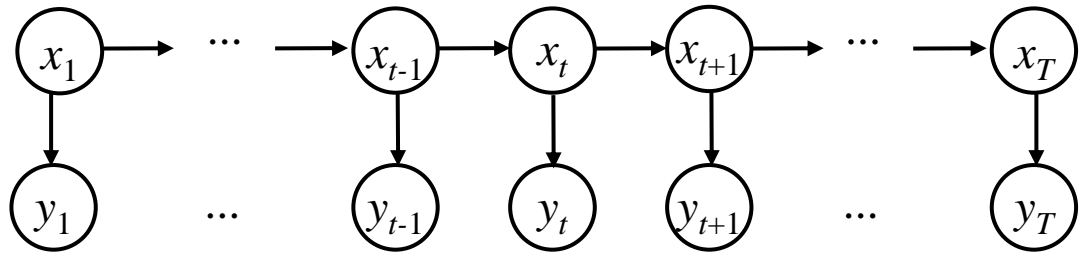
$p(x_t | x_{t-1})$   $p(y_t | x_t)$  局部函数均是典范函数

$$x_t = Ax_{t-1} + G\omega_{t-1} \quad \omega_t \sim N(0, Q)$$

$$y_t = Cx_t + v_t \quad v_t \sim N(0, R)$$

函数的递归计算化归为典范参数的递归计算 ☺

$$p(x_t | y_{1:t}) = ?$$



### 测量更新

$$p(x_t | y_{1:t}) = \frac{p(y_{1:t-1})}{p(y_{1:t})} p(y_t | x_t) p(x_t | y_{1:t-1})$$

$$p(x_1 | y_1) = \frac{1}{p(y_1)} p(y_1 | x_1) p(x_1)$$

$$p(x_2 | y_2) = \frac{p(y_{1,2})}{p(y_1)} p(y_2 | x_2) p(x_2 | y_1)$$

### 时间更新

$$p(x_{t+1} | y_{1:t}) = \int p(x_{t+1} | x_t) p(x_t | y_{1:t}) dx_t$$

$$p(x_2 | y_1) = \int p(x_2 | x_1) p(x_1 | y_1) dx_1$$

$$p(x_1) \xrightarrow{\text{red}} p(x_1 | y_1) \xrightarrow{\text{blue}} p(x_2 | y_1) \xrightarrow{\text{red}} p(x_2 | y_{1:2}) \xrightarrow{\text{dashed red}} p(x_t | y_{1:t})$$

1 个  
典范函数

$M^K$  个  
典范函数  
的混合

$M^K$  个  
典范函数  
的混合

$(M^K)^2$  个  
典范函数  
的混合

$(M^K)^t$  个  
典范函数  
的混合

# Importance Sampling (unnormalized)

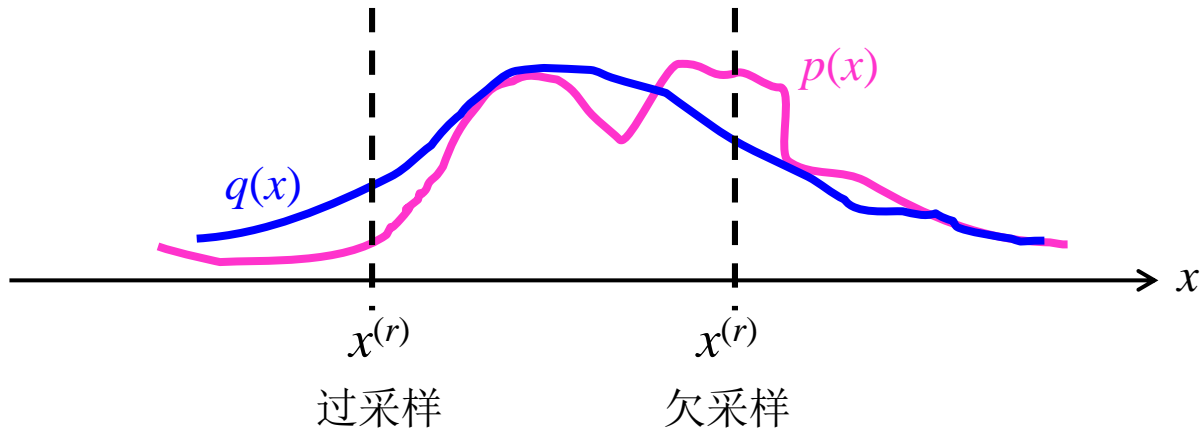
很难从  $p(x)$  — target density 直接采样!

❖ 从另一个分布  $q(x)$  — proposal density 采样, 然后加权

- 随机生成  $N$  个服从分布  $q(x)$  的独立样本  $x^{(r)}$   $w(x)$

$$E_{p(x)}[\phi(x)] = \int \frac{p(x)}{q(x)} \phi(x) q(x) dx = E_{q(x)} \left[ \frac{p(x)}{q(x)} \phi(x) \right]$$
$$\approx \frac{1}{R} \sum_{r=1}^R w(x^{(r)}) \phi(x^{(r)})$$

$x^{(r)}$  连同  $w(x^{(r)})$  视为 概率分布  $p(x)$  的粒子表示



# Importance Sampling (normalized)

❖  $p(x)$  is only known up to a normalization constant  $\alpha$

- $p(x) = p'(x) / \alpha$
- 方便得到与其成正比的  $p'(x)$

$$E_{p(x)}[\phi(x)] = \int \frac{p(x)}{q(x)} \phi(x) q(x) dx = \frac{1}{\alpha} \int \frac{p'(x)}{q(x)} \phi(x) q(x) dx \approx \frac{1}{\alpha} \frac{1}{R} \sum_r w'(x^{(r)}) \phi(x^{(r)})$$

$$\alpha = \int p'(x) dx = \int \frac{p'(x)}{q(x)} q(x) dx \approx \frac{1}{R} \sum_r \frac{p'(x^{(r)})}{q(x^{(r)})} = \frac{1}{R} \sum_r w'(x^{(r)})$$

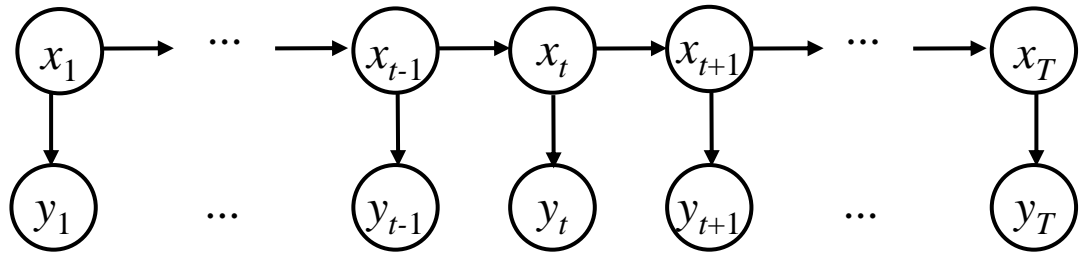
$$E_{p(x)}[\phi(x)] \approx \frac{\sum_r w'(x^{(r)}) \phi(x^{(r)})}{\sum_r w'(x^{(r)})} = \sum_r w^{(r)} \phi(x^{(r)}) \quad w^{(r)} = w(x^{(r)}) = \frac{w'(x^{(r)})}{\sum_r w'(x^{(r)})}$$

$x^{(r)}$  连同  $w^{(r)}$  视为概率分布  $p(x)$  的粒子表示

只要能求出  $\frac{\text{目标分布 } p(x^{(r)})}{\text{举荐分布 } q(x^{(r)})}$  的相对值!



$$p(x_t | y_{1:t}) = ?$$



## 测量更新

$$p(x_t | y_{1:t}) = \frac{p(y_{1:t-1})}{p(y_{1:t})} p(y_t | x_t) p(x_t | y_{1:t-1}) \frac{p(x_t^{(r)} | y_{1:t})}{p(x_t^{(r)} | y_{1:t-1})} \propto p(y_t | x_t^{(r)}) \propto w_t^{(r)}$$

$$p(x_1 | y_1) = \frac{1}{p(y_1)} p(y_1 | x_1) p(x_1) \quad \frac{p(x_1^{(r)} | y_1)}{p(x_1^{(r)})} \propto p(y_1 | x_1^{(r)}) \propto w_1^{(r)}$$

## 时间更新

$$p(x_{t+1} | y_{1:t}) = \int p(x_{t+1} | x_t) p(x_t | y_{1:t}) dx_t = \sum_r w_t^{(r)} p(x_{t+1} | x_t^{(r)})$$

$$p(x_2 | y_1) = \int \frac{p(x_2 | x_1) p(x_1 | y_1)}{\phi(x_1)} dx_1 = \sum_r w_1^{(r)} \frac{p(x_2 | x_1^{(r)})}{\phi(x_1^{(r)})}$$

$$p(x_1) \xrightarrow{\text{red}} p(x_1 | y_1) \xrightarrow{\text{blue}} p(x_2 | y_1) \xrightarrow{\text{red}} p(x_2 | y_{1:2}) \xrightarrow{\text{dashed red}} p(x_t | y_{1:t})$$

样本表示

$$x_1^{(r)}$$

粒子表示

$$x_1^{(r)}, w_1^{(r)}$$

样本表示

$$x_2^{(r)}$$

粒子表示

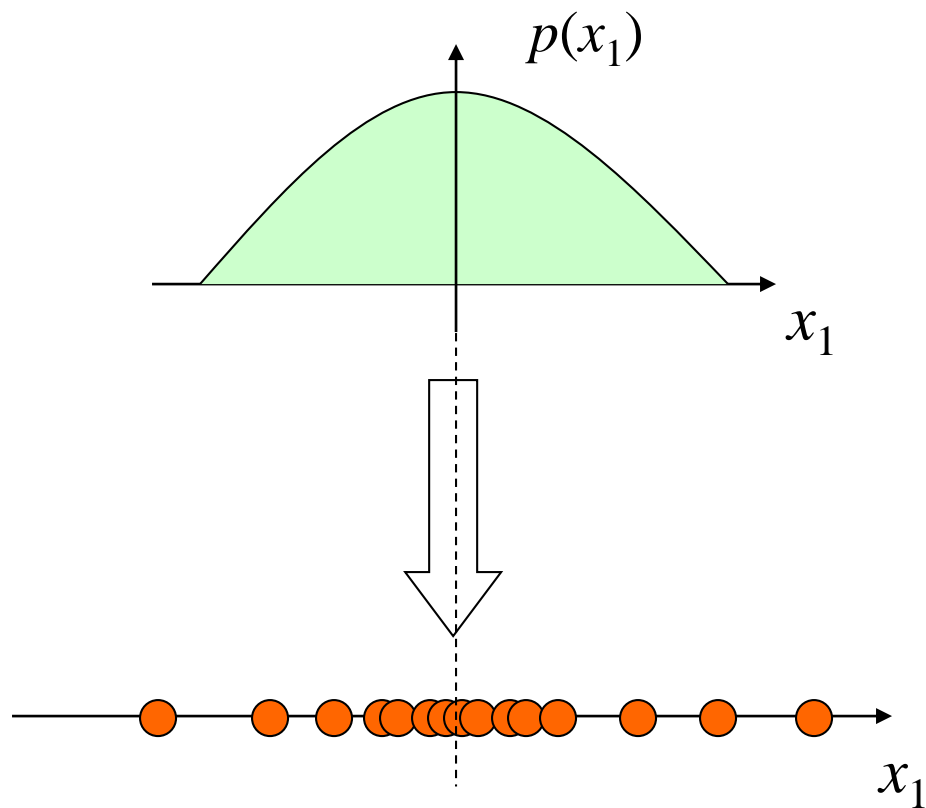
$$x_2^{(r)}, w_2^{(r)}$$

粒子表示

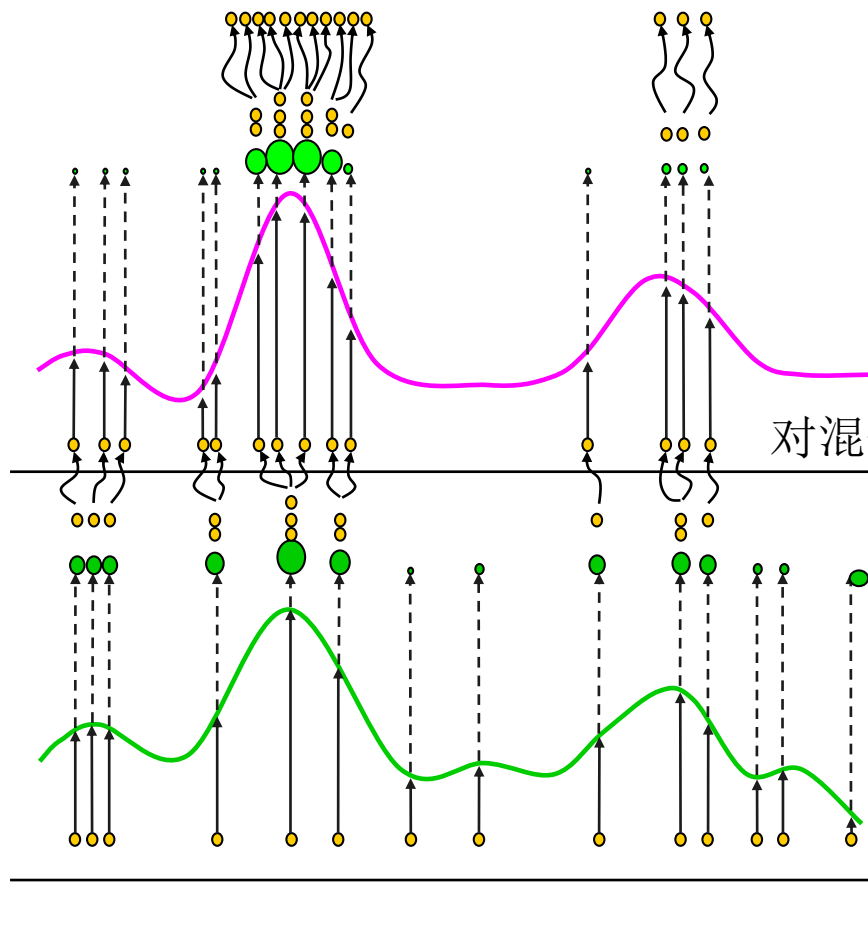
$$x_t^{(r)}, w_t^{(r)}$$

# 初始状态分布

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# Particle filter—直观认识



$$p(y_2 | x_2^{(r)}) \propto w_2^{(r)}$$

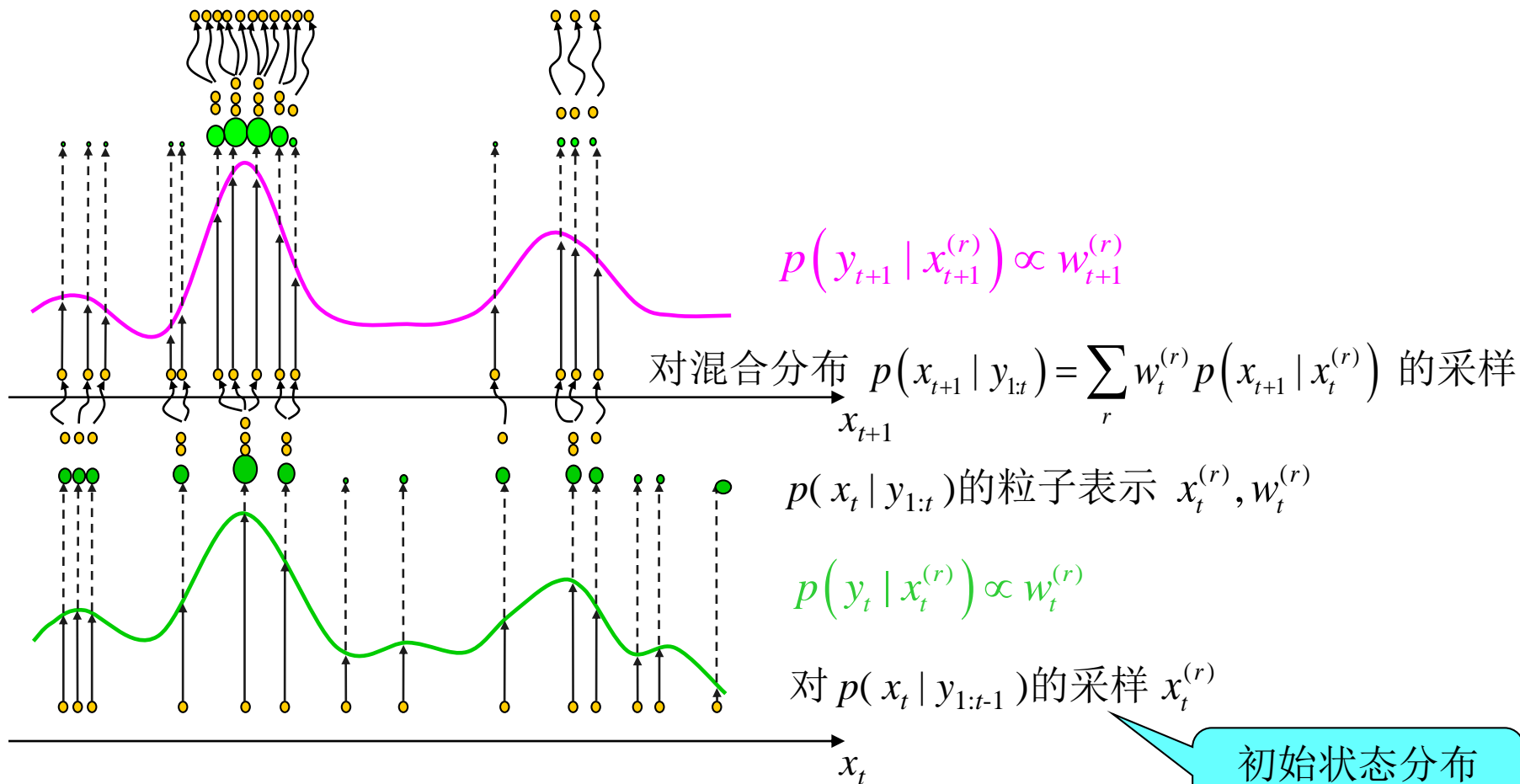
对混合分布  $p(x_2 | y_1) = \sum_r w_1^{(r)} p(x_2 | x_1^{(r)})$  的采样

$p(x_1 | y_1)$  的粒子表示  $x_1^{(r)}, w_1^{(r)}$

$$p(y_1 | x_1^{(r)}) \propto w_1^{(r)}$$

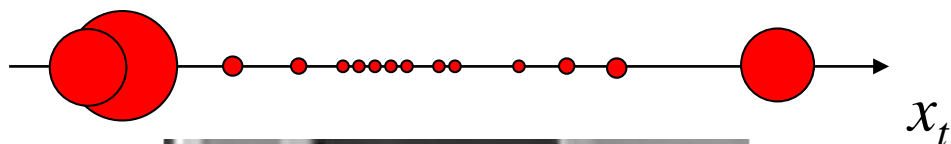
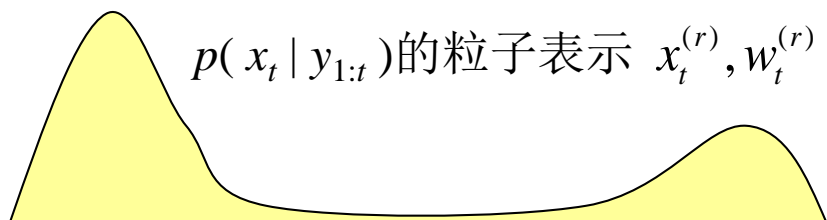
对  $p(x_1)$  的采样  $x_1^{(r)}$

# Particle filter—直观认识



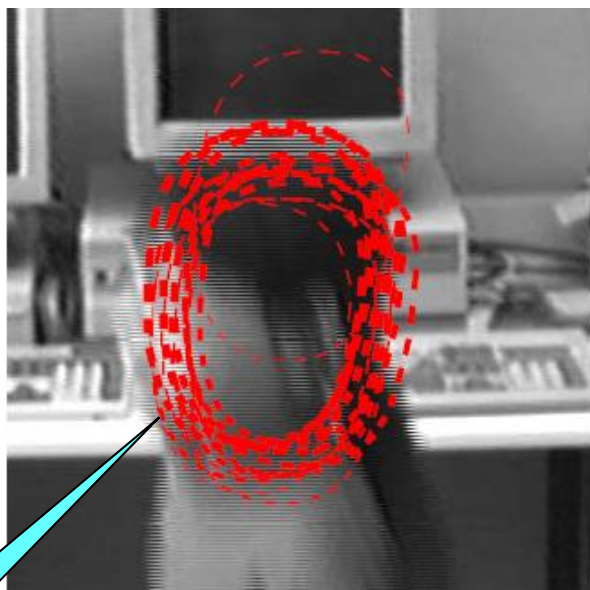
初始状态分布  
 $p(x_1)$

# 联系应用—粒子物理意义



跟踪结果: 条件均值  $E(x_t | y_{1:t})$

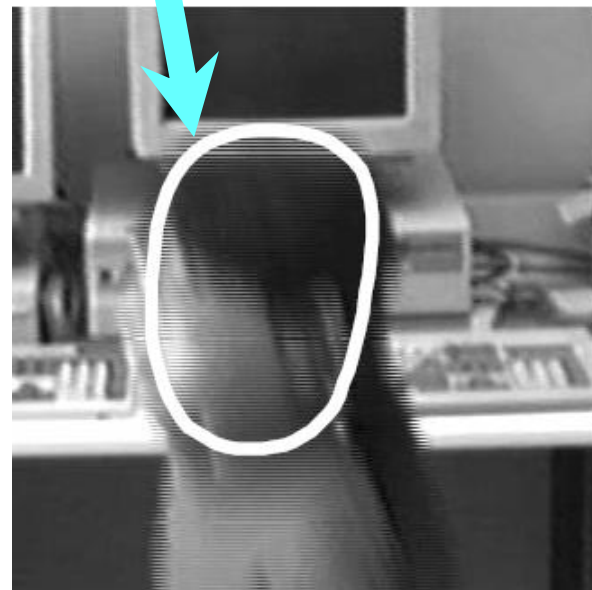
$$= \sum_{r=1}^R w_t^{(r)} \cdot x_t^{(r)}$$



粒子

$x_t^{(r)}$  特定轮廓线

$w_t^{(r)}$  权重(线的粗细)



# Particle filter—演示

- ❖ 基于概率图模型的运动物体跟踪系统
  - 2005年SRT项目校优秀奖
  - 无24 李亚斯等



# 课程章节

- ❖ 第一章 引言 (**1**)
- ❖ 第二章 图模型的表示理论 (**2**)
  - **Semantics (DGM, UGM)**
  - **HMM, CRF**
- ❖ 第三章 图模型的推理理论 (**6**)
  - 精确推理: **variable-elimination, cluster-tree, triangulate**
  - 连续变量: **Kalman**
  - 采样近似: **sampling**
  - 变分近似: **variational**
- ❖ 第四章 图模型的学习理论 (**3**)
  - 参数学习: **maxlikelihoodEstimate, RFLearning, BayesEstimate**
  - 结构学习: **StructureLearning**
- ❖ 第五章 一个综合例子 (**1**)