

概率图模型理论及应用

Theory and Applications of Probabilistic Graphical Models
(Lesson 12 - BayesEstimate)

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课前摘要



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abs_lesson12_BayesEstimate_杨成竹.

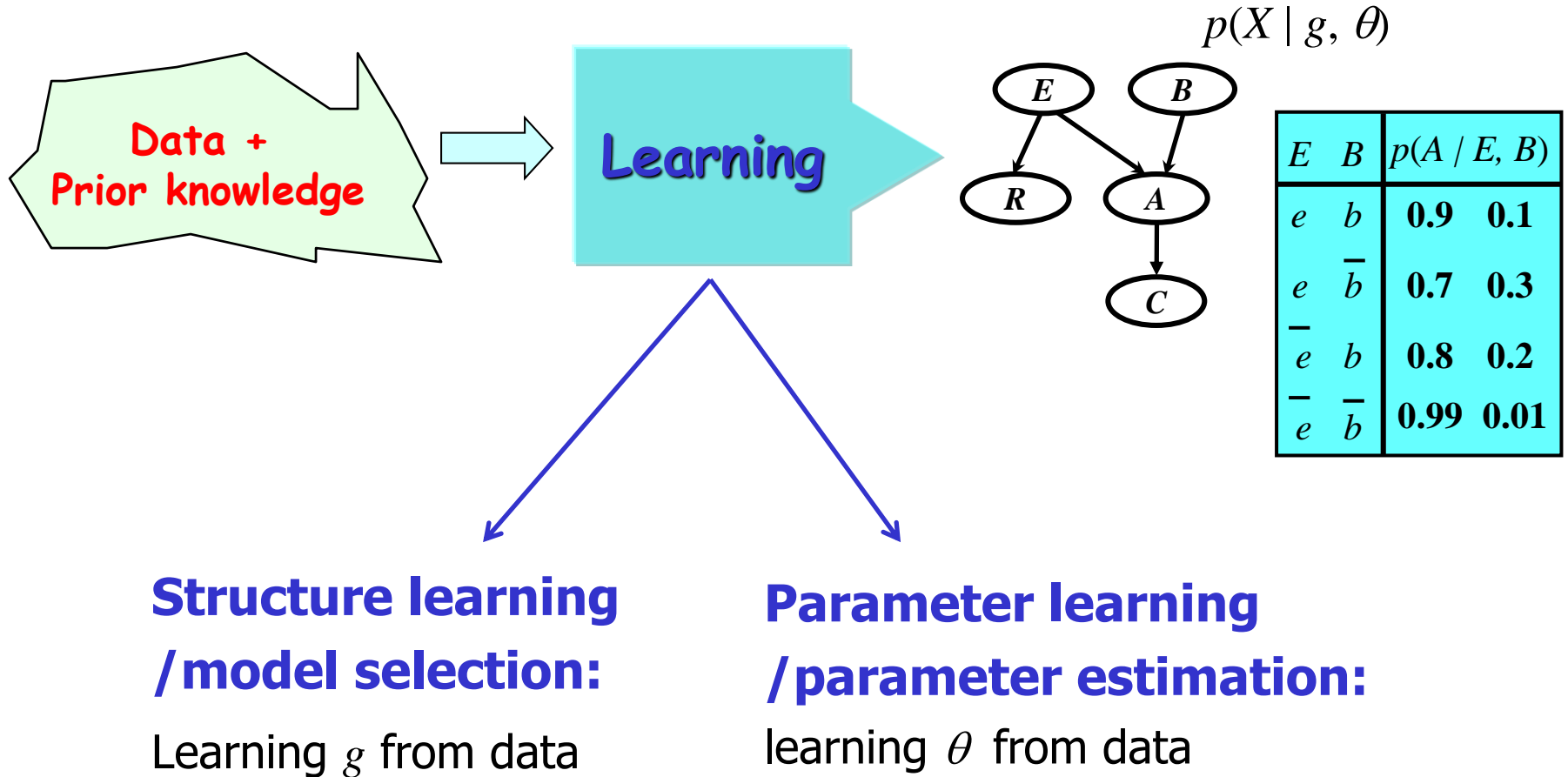


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课程章节

- ❖ 第一章 引言 (**1**)
- ❖ 第二章 图模型的表示理论 (**2**)
 - **Semantics (DGM, UGM)**
 - **HMM, CRF**
- ❖ 第三章 图模型的推理理论 (**6**)
 - 精确推理: **variable-elimination, cluster-tree, triangulate**
 - 连续变量: **Kalman**
 - 采样近似: **sampling**
 - 变分近似: **variational**
- ❖ 第四章 图模型的学习理论 (**3**)
 - 参数学习: **maxlikelihoodEstimate, RFLearning, BayesEstimate**
 - 结构学习: **StructureLearning**
- ❖ 第五章 一个综合例子 (**1**)

Learning



The learning problem

	Known structure	Unknown structure
Complete data	ML	Bayesian
Incomplete data	ML	Bayesian

lesson04_mlEstimate today

Why Bayesian ?

- ❖ 假设抛一块硬币5次，数字朝上($x=0$) 3次
 - 最大似然参数估计 $p(x=0) = 3/5$
 - 这不是一个太好的估计
 - 考虑先验知识——参数 θ 的先验分布 $p(\theta)$
数字朝上的概率应该近似等于0.5
 - 数据稀疏时避免过拟合
 - 贝叶斯方法在结构学习中具有独特的优势
 - Provide uncertainty measure

Parameter learning

— Bayesian (Known structure, complete data)

对单个分布的参数进行估计？

对一个贝叶斯网络的全体参数进行估计？

参数估计的贝叶斯方法

- 给定一个概率分布函数的参数表达式 (parametric form)

$$p(x | \theta)$$

从独立同分布样本集 $D = (x[1], \dots, x[M])$ 中估计出参数 θ ?

- 将 θ 视为一个随机变量

$$p(\theta | D) \propto \frac{p(D | \theta) p(\theta)}{p(D)}$$

- 采取某种准则，从后验分布 $p(\theta | D)$ 出发得到对 θ 的点估计 $\hat{\theta}$

$$\hat{\theta}^{MMSE} = \arg \min_{\hat{\theta}} E \left[\|\hat{\theta} - \theta\|^2 \right] = \int \theta p(\theta | D) d\theta = E(\theta | D)$$

最小均方误差下 贝叶斯估计： 后验均值

$$\hat{\theta}^{MAP} = \arg \max_{\hat{\theta}} E \left[\delta(\theta - \hat{\theta}) \right] = \arg \max_{\hat{\theta}} p(\hat{\theta} | D) = \arg \max_{\hat{\theta}} p(D | \hat{\theta}) p(\hat{\theta})$$

二值相似度下 贝叶斯估计： 最大后验估计

Fully Bayesian

Data: $\mathcal{D} = (x[1], \dots, x[M])$

Prior $P(\theta)$

Posterior $P(\theta|\mathcal{D}) \propto P(\mathcal{D}|\theta)P(\theta)$

Prediction $P(x^{(t)}|\mathcal{D}) = E_{P(\theta|\mathcal{D})}[P(x^{(t)}|\theta)] = \int_{\theta} P(x^{(t)}|\theta)P(\theta|\mathcal{D})d\theta$

$P(\theta|\mathcal{D})$ can be approximated via Markov Chain Monte Carlo methods.

Prediction can be approximated by Monte Carlo averaging:

$$P(x^{(t)}|\mathcal{D}) \approx \frac{1}{n} \sum_{i=1}^n P(x^{(t)}|\theta_i) , \theta_i \sim P(\theta|\mathcal{D})$$

Multinomial distribution 多元分布

- $x \in \{1, 2, \dots, K\}$ is discrete r.v.
- $\theta_k = p(x=k)$, $1 \leq k \leq K$, is the parameters, $\theta = \{\theta_k \mid 1 \leq k \leq K\}$
- 观测到独立同分布样本集 $D = (x[1], \dots, x[M])$
- 希望估计 θ ?



$$\text{似然函数 } p(x[1:M] | \theta) = \prod_{m=1}^M p(x[m] | \theta) = \prod_{k=1}^K \theta_k^{N_k}$$

N_k : 在样本集中 $x[m]=k$ 出现的次数

考虑参数 θ 服从Dirichlet分布

$$p(\theta) = \frac{1}{Z(\alpha)} \prod_{k=1}^K \theta_k^{\alpha_k - 1}$$

$Z(\alpha)$ 是归一化常数, $\alpha = (\alpha_1, \dots, \alpha_K)$ 称为hyperparameters

Dirichlet分布 $p(\theta) = \frac{1}{Z(\alpha)} \prod_{k=1}^K \theta_k^{\alpha_k - 1}$

❖ $Z(\alpha)$ 是归一化常数

$$Z(\alpha) = \int_{\theta_1} \cdots \int_{\theta_K} \theta_1^{\alpha_1 - 1} \cdots \theta_K^{\alpha_K - 1} d\theta_1 \cdots d\theta_K = \frac{\Gamma(\alpha_1) \cdots \Gamma(\alpha_K)}{\Gamma(\alpha_1 + \cdots + \alpha_K)}$$

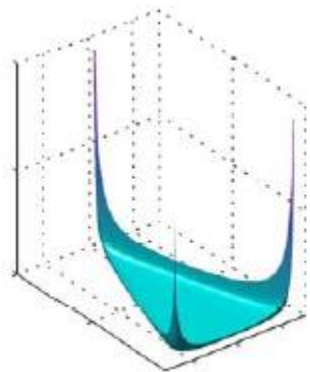
- $\Gamma(\alpha)$ is gamma function: $\Gamma(\alpha) = \int_0^{\infty} t^{\alpha-1} e^{-t} dt$
- For integers, $\Gamma(n+1) = n!$

❖ 如果 $\theta = (\theta_1, \dots, \theta_K) \sim \text{Dirichlet}(\alpha_1, \dots, \alpha_K)$

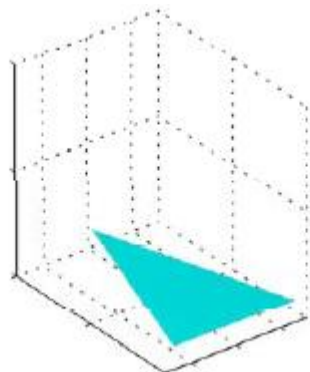
$$E[\theta_k] = \int \theta_k \cdot p(\theta) d\theta = \frac{\alpha_k}{\sum_{\ell} \alpha_{\ell}}$$

Dirichlet分布 (K=3) $p(\theta) = \frac{1}{Z(\alpha)} \theta_1^{\alpha_1-1} \theta_2^{\alpha_2-1} \theta_3^{\alpha_3-1}$

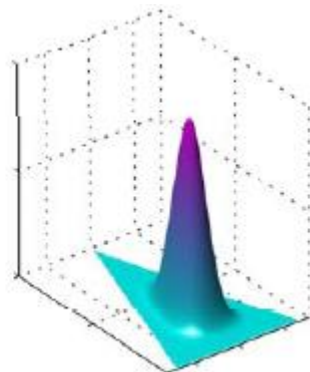
- ❖ $\theta = (\theta_1, \dots, \theta_K) \sim \text{Dirichlet}(\alpha_1, \dots, \alpha_K)$ $\sum_k \theta_k = 1, \theta_k \geq 0$
 θ 定义在维数为 $K-1$ 的单纯形上



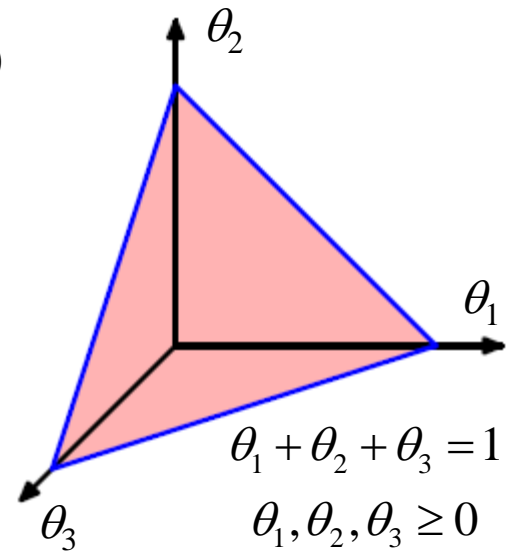
$$\alpha_1 = \alpha_2 = \alpha_3 = 0.1$$



$$\alpha_1 = \alpha_2 = \alpha_3 = 1$$



$$\alpha_1 = \alpha_2 = \alpha_3 = 10$$



Dirichlet后验分布 \propto 多元分布似然函数 \cdot Dirichlet先验分布

$$p(\theta | D) \propto p(D | \theta) p(\theta)$$

- The likelihood function $p(D | \theta) = \prod_{k=1}^K \theta_k^{N_k}$
- The Dirichlet prior $p(\theta) \propto \prod_{k=1}^K \theta_k^{\alpha_k - 1}$
- The posterior probability θ of given $D \sim \text{Dirichlet}(\alpha_1 + N_1, \dots, \alpha_K + N_K)$

$$p(\theta | D) \propto \prod_{k=1}^K \theta_k^{N_k} \times \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{\alpha_k + N_k - 1}$$

Dirichlet is the conjugate prior for multinomial

- (最小均方误差下)贝叶斯估计

$$\hat{\theta}_k^{MMSE} = \frac{\alpha_k + N_k}{\sum_l (\alpha_l + N_l)} \quad \hat{\theta}_k^{ML} = \frac{N_k}{\sum_{l=1}^K N_l}$$

超参数 $\alpha_1, \dots, \alpha_K$ 可视为一种根据先验知识而设定的先验次数 (prior count)

Plate notation

- ❖ 服从总体分布 $p(x | \theta)$, 独立同分布采样 $D = (x[1], \dots, x[M])$ 的贝叶斯网络表示

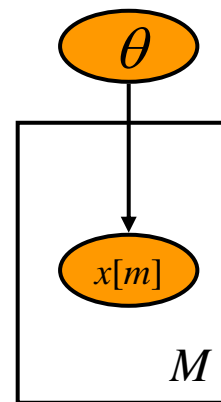
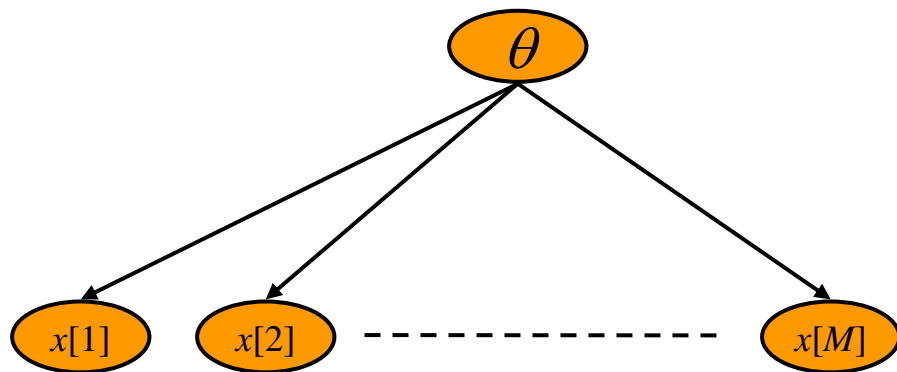


Plate notation

- 用于表示重复结构 (repetitive structure)

1) 将盒子内的图结构重复多次。

重复的次数由右下角的数字 (e.g. M) 来指定,
盒子内变量的编号相应变动 (e.g. m)

2) 进入盒子、离开盒子的有向边 进行相应的重复。

Learning parameters for BNs (complete data)

- 考虑贝叶斯网络 $x = \{x_1, \dots, x_N\}$

假设：各个条件分布 $p(x_1|pa_1), \dots, p(x_N|pa_N)$ 有各自表征参数 $\{\theta_1, \dots, \theta_N\}$

头姿类别 $x_1 \in 1:K$

观测图像 $x_2 \in R^{44*28}$

总体分布： $p(x_1, x_2/\theta)$

$p(x_2|x_1, \theta_2)$ 的参数： $\{\mu_k, \Sigma_k\}_{k=1:K}$

$p(x_1|\theta_1)$ 的参数： $\{\pi_k\}_{k=1:K}$

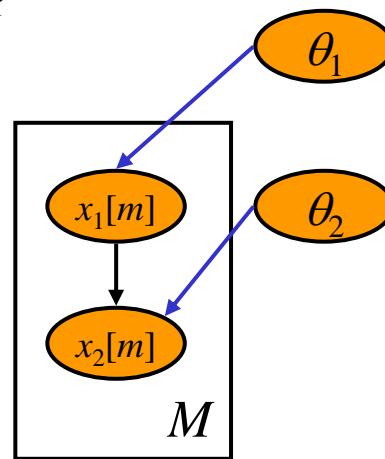
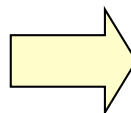
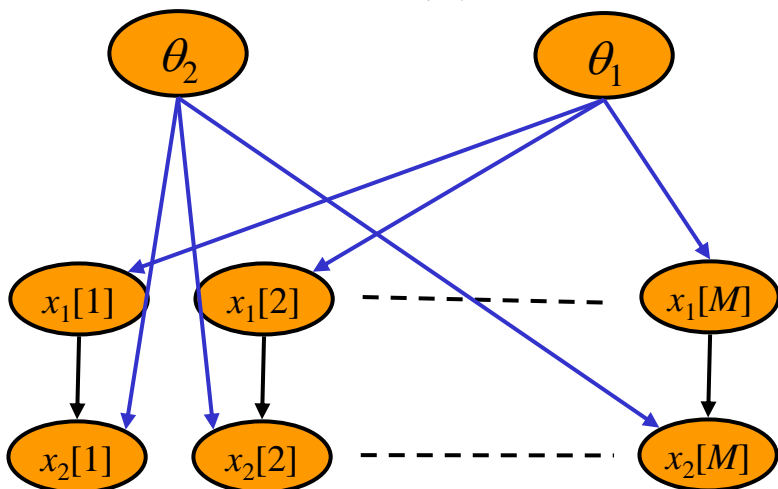
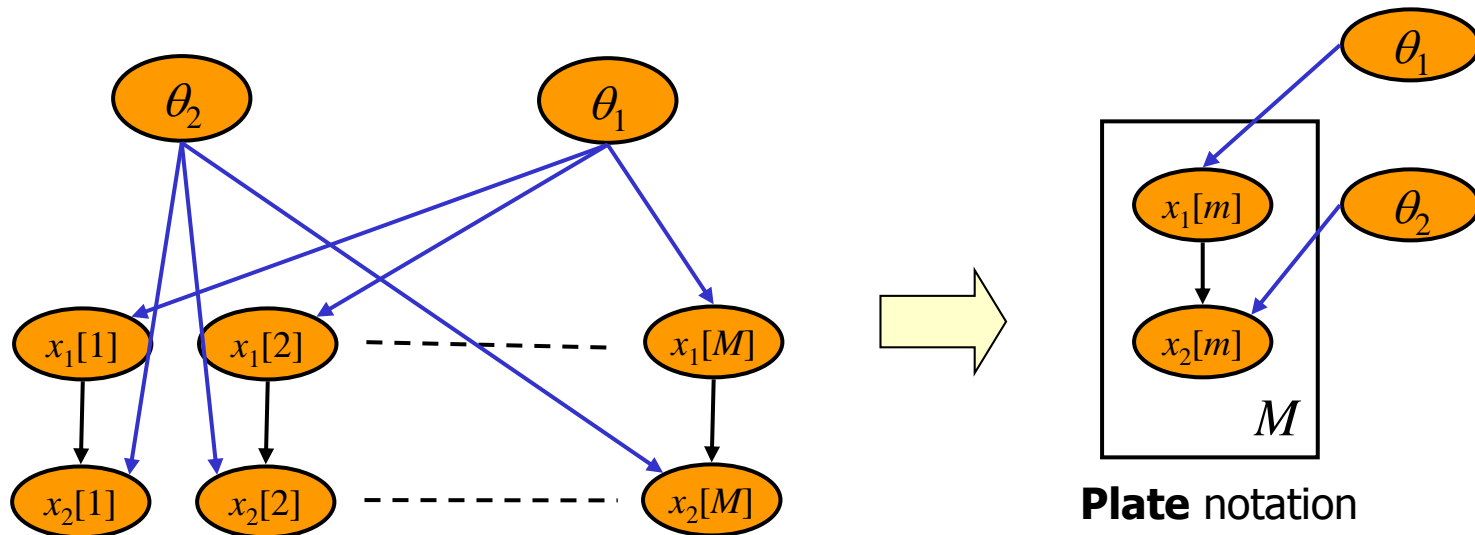


Plate notation

Learning parameters for BNs (complete data)



- Definition: Global parameter independence $p(\theta) = \prod_{n=1}^N p(\theta_n)$

$$p(\theta | D) \propto p(\theta) \times p(D | \theta) = p(\theta) \times \prod_{m=1}^M p(x[m] | \theta)$$

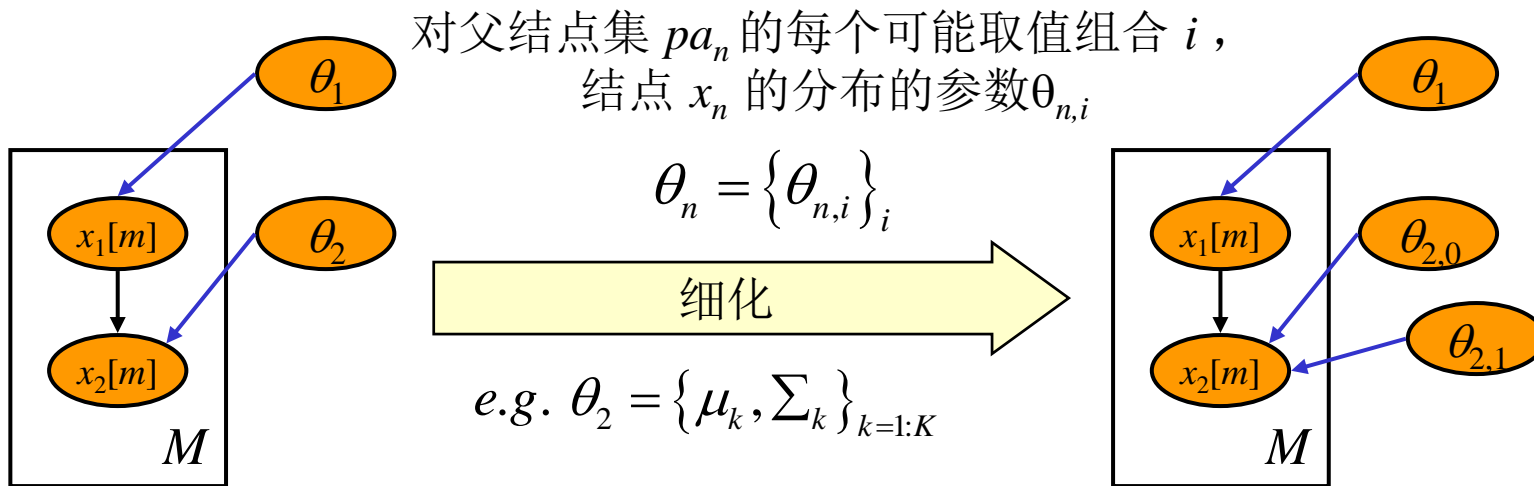
$$= p(\theta) \times \prod_{m=1}^M \prod_{n=1}^N p(x_n[m] | pa_n[m], \theta_n)$$

$$p(\theta | D) \propto \prod_{n=1}^N \left\{ p(\theta_n) \times \prod_{m=1}^M p(x_n[m] | pa_n[m], \theta_n) \right\}$$

$$p(\theta | D) = \prod_{n=1}^N p(\theta_n | D)$$

总体参数的后验分布
= 每个结点处表征参数的后验分布的连乘积

Learning parameters for BNs (complete data)



e.g. $p(\{\mu_k, \Sigma_k\}_{k=1:K}) = \prod_k p(\mu_k, \Sigma_k)$

- Definition: Local parameter independence $p(\theta_n) = \prod_i p(\theta_{n,i})$

$$p(\theta_n | D) \propto p(\theta_n) \times \prod_{m=1}^M p(x_n[m] | pa_n[m], \theta_n)$$

$$= \prod_i \left\{ p(\theta_{n,i}) \cdot \prod_{\substack{1 \leq m \leq M \\ \text{s.t. } pa_n[m]=i}} p(x_n[m] | pa_n[m]=i, \theta_{n,i}) \right\}$$

$$p(\theta_n | D) = \prod_i p(\theta_{n,i} | D)$$

Bayes estimate for multinomial Bayes net

对每个结点 n 及父结点集 pa_n 的每个可能取值组合 i
一个单独的多元分布的参数 $\theta_{n,i,k}$

- 充分统计量：次数

$$N_{n,i,k} = \sum_{m=1}^M \delta(pa_n[m] = i, x_n[m] = k)$$

$$\left[\hat{\theta}_{n,i,k} \right]^{ML} = \frac{N_{n,i,k}}{\sum_{l=1}^{K_n} N_{n,i,l}}$$

- 假设局部参数均服从Dirichlet分布 $p(\theta_{n,i}) \sim \text{Dirichlet}(\alpha_{n,i,1}, \alpha_{n,i,2}, \dots, \alpha_{n,i,K_n})$

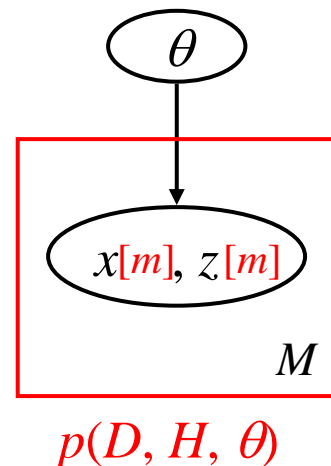
$$\left[\hat{\theta}_{n,i,k} \right]^{MMSE} = \frac{\alpha_{n,i,k} + N_{n,i,k}}{\sum_{l=1}^{K_n} (\alpha_{n,i,l} + N_{n,i,l})}$$

Parameter learning

— Bayesian (Known structure, incomplete data)

一般原理

- ❖ 总体分布 $p(x, z | \theta)$ ，参数先验分布 $p(\theta)$
 - 总体分布的IID 样本集 $D = (x[1], \dots, x[M])$ ：观测数据
 $H = (z[1], \dots, z[M])$
- ❖ 观测到数据 D ，求参数的后验分布 $p(\theta | D)$?
 - 参数 θ 视为一种特殊的隐变量，化归为推理计算
- ❖ 变分贝叶斯方法 (Variational Bayesian)
 - 基于变分推理 求解 参数后验分布 $p(\theta | D)$
 - 用变分分布 $q(H, \theta)$ 去近似真实后验分布 $p(H, \theta | D)$

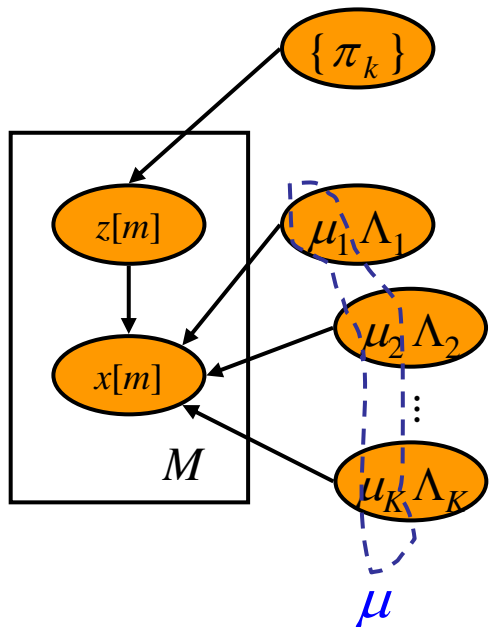


$$q(H, \theta) = \underbrace{q(H)}_{p(H|D)} \underbrace{q(\theta)}_{p(\theta|D)}$$

轮换求解：

- $\log q(\theta) = E_{q(H)} [\log p(H, D, \theta) | \theta] + const = \sum_H q(H) \log p(H, D, \theta) + const$
- $\log q(H) = E_q [\log p(H, D, \theta) | H] + const = \sum_\theta q(\theta) \log p(H, D, \theta) + const$

高斯混合模型参数估计的贝叶斯方法



总体分布 $p(z = k, x) = p(z = k) p(x | z = k)$
 $= \pi_k N(x | \mu_k, \Lambda_k)$

参数先验分布 $p(\theta) = p(\pi) \prod_{k=1}^K p(\mu_k, \Lambda_k)$

Dirichlet prior for mixing coefficients

$$p(\pi) = C(\alpha_0) \prod_{k=1}^K \pi_k^{\alpha_0 - 1}$$

Normal-Wishart prior for means and precisions

$$p(\mu_k, \Lambda_k) = \mathcal{N}(\mu_k | \mathbf{m}_0, (\beta_0 \Lambda_k)^{-1}) \mathcal{W}(\Lambda_k | \mathbf{W}_0, \nu_0)$$

高斯混合模型参数估计的贝叶斯方法

- ❖ 假设如下的变分分布

$$q(z[1:M], \pi, \mu, \Lambda) = q(z[1:M])q(\pi, \mu, \Lambda)$$

$$q(H, \theta) = q(H)q(\theta)$$

No other assumptions!

- ❖ 变分推理结果



$$q(\pi, \mu, \Lambda) = q(\pi) \prod_{k=1}^K q(\mu_k, \Lambda_k)$$

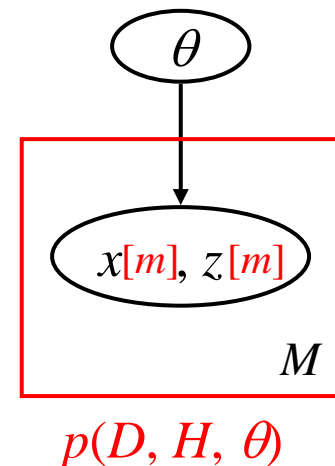
~ Dirichlet (above $q(\pi)$)
~ Normal-Wishart (above $q(\mu_k, \Lambda_k)$)

$$q(z[1:M]) = \prod_{m=1}^M q(z[m]) \sim \text{Multinomial}$$

~ Multinomial (to the right of the product)

详见Bishop书 10.2 Illustration: Variational Mixture of Gaussians

VB discussion: 点估计



❖ 总体分布 $p(x, z | \theta)$ ，参数先验分布 $p(\theta)$

- 总体分布的IID 样本集 $D = (x[1], \dots, x[M])$ ：观测数据
 $H = (z[1], \dots, z[M])$

❖ 观测到数据 D ，求参数的后验分布 $p(\theta | D)$ ？

❖ 变分贝叶斯方法（Variational Bayesian）

$$p(H, \theta | D) \approx q(H, \theta) = q(H)q(\theta)$$

轮换求解：

- $\log q(\theta) = E_q[\log p(H, D, \theta) | \theta] + const = \sum_H q(H) \log p(H, D, \theta) + const$

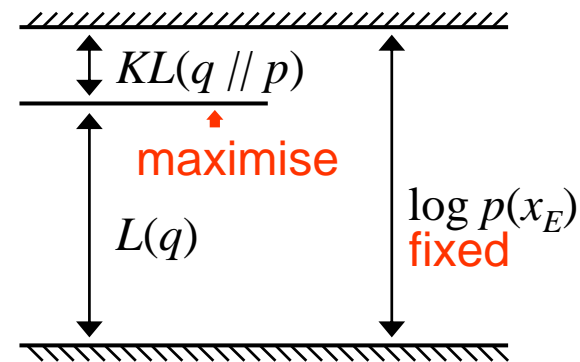
- $\log q(H) = E_q[\log p(H, D, \theta) | H] + const = \sum_{\theta} q(\theta) \log p(H, D, \theta) + const$

只关心 θ 的点估计 $\xrightarrow[\text{约束 } q(\theta) = \delta(\theta - \theta^*)]{\text{约束}}$ $\min_{\text{约束 } q(\theta) \text{ 为 } \delta \text{ 函数}} KL[q(H)q(\theta) \| p(H, \theta | D)]$

VB discussion: 点估计

❖ 变分贝叶斯方法 (Variational Bayesian)

$$\min_{\text{约束 } q(x_k) \text{ 为 } \delta \text{ 函数}} KL \left[\prod_{i \in H} q(x_i) \parallel p(x_H | x_E) \right]$$



针对 $q(x_k)$ 的最优化: 将 $L(q)$ 视为 $q(x_k)$ 的函数, 与 $q(x_k)$ 无关项并入常数

$$\begin{aligned} L(q) &= H(q(x_k)) + \sum_{x_k} q(x_k) \log \tilde{p}(x_k | x_E) + \text{常数} \\ &= -KL(q(x_k) \parallel \tilde{p}(x_k | x_E)) + \text{常数} \end{aligned}$$

无约束最优化: $\log q(x_k) = \log \tilde{p}(x_k | x_E) = E_q[\log p(x_H, x_E) | x_k] + \text{常数}$

有约束最优化: $q(x_k) = \delta(x_k - x_k^*)$, 其中 $x_k^* = \arg \max_{x_k} \log \tilde{p}(x_k | x_E)$

约束 $q(x_k)$ 为 δ 函数

$$= \arg \max_{x_k} E_q[\log p(x_H, x_E) | x_k]$$

From VB to 不完备数据下MAP估计

$$q(H, \theta) = q(H)q(\theta)$$

轮换求解:

■ $\log q(\theta) = E_q [\log p(H, D, \theta) | \theta] + const$ $x_k^* = \arg \max E_q [\log p(x_H, x_E) | x_k]$

只关心 θ 的点估计 $\xrightarrow[\text{约束 } q(\theta) = \delta(\theta - \theta^*)]{}$ $\theta^* = \arg \max_{\theta} E_q [\log p(H, D, \theta) | \theta]$

$$\theta^* = \arg \max_{\theta} \sum_H p(H | D, \theta^{(old)}) \log p(H, D, \theta)$$

$$\theta^* = \arg \max_{\theta} \left\{ \sum_H p(H | D, \theta^{(old)}) \log p(H, D | \theta) + \log p(\theta) \right\}$$

■ $\log q(H) = E_q [\log p(H, D, \theta) | H] + const$

$$\begin{aligned} \log q(H) &= \sum_{\theta} q(\theta) \log p(H, D, \theta) + const \\ &= \log p(H, D, \theta^{(old)}) + const \end{aligned}$$

$$q(H) \propto p(H | D, \theta^{(old)})$$

ML: $\max_{\theta} p(D | \theta)$
 MAP: $\max_{\theta} p(D | \theta) p(\theta)$

From VB to ICM (Iterative conditional modes)

$$q(H, \theta) = q(H)q(\theta) \quad \text{众数, 最频值, 最常出现的变量值}$$

轮换求解:

$$\blacksquare \log q(\theta) = E_q [\log p(H, D, \theta) | \theta] + \text{const}$$

$$\text{只关心 } \theta \text{ 的点估计} \xrightarrow[q(\theta) = \delta(\theta - \theta^*)]{\text{约束}} \theta^* = \arg \max_{\theta} E_q [\log p(H, D, \theta) | \theta]$$
$$= \sum_H q(H) \log p(H, D, \theta)$$

$$\theta^* = \arg \max_{\theta} \log p(\theta | D, H^*)$$

$$\blacksquare \log q(H) = E_q [\log p(H, D, \theta) | H] + \text{const}$$

$$\text{只关心 } H \text{ 的点估计} \xrightarrow[q(H) = \delta(H - H^*)]{\text{约束}} H^* = \arg \max_H E_q [\log p(H, D, \theta) | H]$$
$$= \sum_{\theta} q(\theta) \log p(H, D, \theta)$$

$$H^* = \arg \max_H \log p(H | D, \theta^*)$$

推理计算: $p(x_H | x_E)$

Gibbs采样:

$$\hat{x}_k - \text{sampling from } p(x_k | x_{H \setminus \{k\}}, x_E)$$

$$\hat{x}_k - \text{sampling from } p(x_H, x_E)$$

均值场变分:

$$\log \hat{q}(x_k) = E_q [\log p(x_H, x_E) | x_k] + \text{const}$$

$$\text{ICM: } \max_{x_H} p(x_H | x_E)$$

$$\min_{\text{约束 } q(x_H) \text{ 为 } \delta \text{ 函数}} KL[q(x_H) \| p(x_H | x_E)]$$

$$x_k^* = \arg \max_{x_k} p(x_H, x_E)$$

轮换采样:

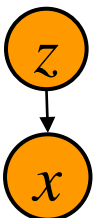
$$\begin{array}{c} \underline{x_1}, x_2, x_3, \dots, x_K \\ \downarrow \\ \hat{x}_1, \underline{x_2}, x_3, \dots, x_K \\ \downarrow \\ x_1, \hat{x}_2, x_3, \dots, x_K \end{array}$$

轮换求期望:

$$\begin{array}{c} \underline{q(x_1)}, q(x_2), q(x_3), \dots, q(x_K) \\ \downarrow \\ \hat{q}(x_1), \underline{q(x_2)}, q(x_3), \dots, q(x_K) \\ \downarrow \\ \hat{q}(x_1), \hat{q}(x_2), q(x_3), \dots, q(x_K) \end{array}$$

轮换求最大化:

$$\begin{array}{c} \underline{x_1}, x_2, x_3, \dots, x_K \\ \downarrow \\ x_1^*, \underline{x_2}, x_3, \dots, x_K \\ \downarrow \\ x_1, \hat{x}_2^*, x_3, \dots, x_K \end{array}$$



Iterative learning

tradeoff efficiency and accuracy

No

θ 是否采取点估计

Yes

No

H 是否采取点估计

Yes

$q(H), q(\theta)$	$q(H), \theta^*$
Expectation-Expectation (EE)	Expectation-Maximization (EM)
Variational Bayes (VB)	Mixture of Gaussians
VB Mixtures of Gaussians	
$H^*, q(\theta)$	H^*, θ^*
Maximization-Expectation (ME)	Maximization-Maximization (MM)
Bayesian K-Means	Iterative Conditional Modes (ICM)
	K-Means

Max Welling and Kenichi Kurihara. Bayesian K-Means as a "Maximization-Expectation" Algorithm. SIAM Conference on Data Mining (SDM2006)

<http://www.ics.uci.edu/~welling/publications/publications.html>

A Comparison of Algorithms for Inference and Learning in Probabilistic Graphical Models

Brendan J. Frey, *Senior Member, IEEE*, and Nebojsa Jojic

The learning problem

	Known structure	Unknown structure
Complete data	ML	Bayesian
Incomplete data	ML	Bayesian

lesson06_mlEstimate today

课程章节

- ❖ 第一章 引言 (**1**)
- ❖ 第二章 图模型的表示理论 (**2**)
 - **Semantics (DGM, UGM)**
 - **HMM, CRF**
- ❖ 第三章 图模型的推理理论 (**6**)
 - 精确推理: **variable-elimination, cluster-tree, triangulate**
 - 连续变量: **Kalman**
 - 采样近似: **sampling**
 - 变分近似: **variational**
- ❖ 第四章 图模型的学习理论 (**3**)
 - 参数学习: **maxlikelihoodEstimate, RFLearning, BayesEstimate**
 - 结构学习: **StructureLearning**
- ❖ 第五章 一个综合例子 (**1**)